the samples in each bracket were independently obtained, an estimate of variance of \hat{p} turns out to be

(2)
$$v(\hat{p}) = \sum_{h=1}^{10} \frac{N_h^2}{N^2} \cdot \frac{N_h - n_h}{N_h} \cdot \frac{\hat{p}_h(1 - \hat{p}_h)}{n_{h-1}}$$

(1) is found to be .6004 with estimated variance (2) computed at .000005479. It may be noted that (1) and (2) are unbiased estimators of their respective population values if the Manila list is a stratified simple random sample of the total 1960 taxable returns in the Philippines.

2.2A Estimating the number or proportion of married taxable returns in sub-brackets

Since the sample sizes of single taxable returns and of married taxable returns are each pre-determined, the number or proportion of married taxable returns in the sub-brackets of the total 1960 Philippine list cannot be estimated using an estimator analogous to (1). An estimator \hat{N}_{m}^{*} , however, can be devised utilizing the per cent distribution of married taxable returns (or of single taxable returns if we use an estimator \hat{N}_{s}^{*} , the subscript s is for single) both in sub-brackets of Table 2.3A and in brackets of Table 2.2, column (5) (or column (3) for single).

Define the following symbols in the Philippine list by: N_{m} = number of married taxable returns

p_m = proportion of married taxable returns

N_{mh} = number of married taxable returns in the hth bracket

δ mh = per cent distribution of the hth bracket among the married taxable returns

N_{mhi} = number of married taxable returns in the ith subbracket of the hth bracket

 δ_{mhi} = per cent distribution of the ith subbracket in the hth bracket of the married taxable returns.

The analogous symbols for the properties in the Manila list and the subsamples are indicated by a circumflex "^" above the symbols. From the definitions $\sum_{i} \delta_{mhi}$ and $\sum_{i} \delta_{mh}$ are each equal to unity. Similarly, $\sum_{i} N_{mhi} = \sum_{i} N_{mh} = N_{m}$. Since $\delta_{mh} = N_{mh}/N_{m}$, $\delta_{mhi} = N_{mhi}/N_{mh}$, and $p_{m} = N_{m}/N$

(1) $p_m \cdot \delta_{mh} \cdot \delta_{mhi} = N_m/N \cdot N_{mh}/N_m \cdot N_{mhi}/N_{mh}$ $= N_{mhi}/N = \delta'_{mhi}, \text{ per cent distribution}$ of the ith subbracket in the hth
bracket among the total list of
1960 taxable returns.

From the last set/equality in (1), $N_{mhi} = N\delta_{mhi}^{\dagger}$. An estimate of N_{mhi} from the sample may be given by

(2)
$$\hat{N}_{mhi} = N\hat{\delta}_{mhi} = N\hat{p}_{m} \cdot \hat{\delta}_{mh} \hat{\delta}_{mhi}$$

Section 2.2 gives $\hat{p}_m = \hat{p} = .6004$. For the second subbracket of the first bracket, $\hat{\delta}_{m1} = .4411$ (Table 2.2, column (5)) and $\hat{\delta}_{m12} = .4233$ (Table 2.3A, under column "married"). From (2) above,

$$\hat{\delta}_{mhi}' = \hat{N}_{mhi}/N = \hat{p}_{m} \cdot \hat{\delta}_{mh} \hat{\delta}_{mhi}$$

$$= (.6004)(.4411)(.4233) = .1121$$

The "1-2" thousand subbracket of the married group amounts to about 11.21% of the total number of taxable returns in 1960. Using (2) above, :an estimated number runs to 103,137 x .1121 = 11,562 married taxable returns in 1960.

To find the variance of \hat{N}_{mhi} , we have

(3)
$$V(\hat{N}_{mhi}) = E[\hat{N}_{mhi} - E(\hat{N}_{mhi})]^{2}$$
$$= E(\hat{N}_{mhi})^{2} - \{E(\hat{N}_{mhi})\}^{2}$$

(4)
$$E(\hat{N}_{mhi}) = E_3 E_2 E_1 (\hat{p}_m \hat{\delta}_{mh} \hat{\delta}_{mhi}),$$

where E₁ refers to conditional expectation over i for fixed h among married taxable returns.

E₂ refers to conditional expectation over h among married taxable returns, and

E₃ refers to expectation of proportion of married taxable returns.

Thus,

(5)
$$E(\hat{N}_{mhi}) = E_{3}[\hat{p}_{m} \cdot E_{2}\{\delta_{mh} \cdot E_{1}(\delta_{mhi})\}]$$

$$= E_{3}[\hat{p}_{m} \cdot E_{2}\{\delta_{mh} \cdot \overline{\delta}_{mh}\}]$$

$$= Np_{m} \text{ cov } (\delta_{mh}, \overline{\delta}_{mh})$$

$$E(\hat{N}_{mhi})^{2} = N^{2} E_{3}[\hat{p}_{m}^{2} \cdot E_{2}\{\delta_{mh}^{2} \cdot E_{1}(\delta_{mhi}^{2})\}]$$

Combining (5) and (6), we obtain an expression for the variance (3). An estimate of (3) can be obtained from the sample by a tedious computational scheme.

2.3 The Projection of Individual Income Tax

The general procedure having been followed was along the same lines as that in Part I; namely, to have an estimate of net taxable income by bracket and an estimate of number of taxable returns for the same bracket. However, it is believed that estimation in this sector has been more difficult than the corporation sector since: (i) the number of brackets involved are more numerous (23 brackets against 4 previously), (ii) certain changes in personal deductions left us only with years 1959 through 1962 which have similar historical exemptions, (iii) the dearth of data on percentage distribution by bracket led us to utilize the percent distribution of the 1960 list for separating the number of taxable returns into taxable brackets.

taxable individual returns is a random sample from an infinite population (that is, the 1960 list is a sample from what could have been all possible lists through time) then the means of taxable assessments for each bracket in the 1960 list can be utilized as estimates of the population means. The population variance and, hence, the population standard deviation may be estimated from the sample of size 1,073 previously obtained. Assuming now the sample variances to be approximately those of the corresponding population values, by the Central Limit Theorem, the hth bracket mean of the 1960 list approaches that of the normal distribution with the same mean and variance equal to $1/n_h$ of the population values, where n_h is the sample size of the hth stratum.

with the 1960 means in each bracket used as the middle estimate, low and high estimates similarly can be obtained from the fiducial limits of the 95% confidence interval estimates. These estimates of the bracket means are tabulated under Table 2.4.

(B) The growth of individual taxable returns from 1959 (t=0) through 1962 were:

t: 0 1 2 3

R_t: 90,364 103,337 118,844 137,935

Transforming R_{t} into logarithms (base 10) and fitting by least squares gives

(3)
$$\log_{10} R_t = 0.78576 + 0.04917t$$

with estimated standard deviation $\hat{\sigma}_{\log_{10}R}$ of \log_{10}^{R} ,

given by

(4)
$$\hat{\sigma}_{\log_{10} R} = .01758.$$

Projections to 1963 and 1964 are obtained by putting t=4 and t=5, respectively, in equation (3). Thus, after taking the antilogarithms of the projected $\log_{10}R_{+}$, we have

Percentage distribution of the 1960 list, when applied to the projected values for 1963 and 1964, respectively, gives results in Table 2.5 below.

Table 2.4 ×

9510 C

INTERVAL ESTIMATES OF MEAN NET INCOME BY TAX BRACKET, FOR PRESENT BRACKET SIZES (Pesos)

| Present | : Mean Net Income | | | | | |
|-------------------|-------------------|----------|-----------|--|--|--|
| Tax Brackets | : Low | : Middle | : High | | | |
| Z_ro- 2,000 | 627 | 697 | 767 | | | |
| 2,001- 4,000 | 2,867 | 2,969 | 3,071 | | | |
| 4,001- 6,000 | 5, 497 | 5,633 | 5,769 | | | |
| 6,001- 8,000 | 7,321 | 7,464 | 7,607 | | | |
| 8,001-10,000 | 9,524 | 9,699 | 9,874 | | | |
| 10,001-20,000 | 13,063 | 13,999 | 14, 935 | | | |
| 20, 001- 30, 000 | > 23,893 | 24, 784 | 25,675 | | | |
| 30,001-40,000 | 33,660 | 34, 566 | 35, 472 | | | |
| 40,001- 50,000 | 45, 854 | 46,689 | 47,524 | | | |
| 50,001-60,000 | 54,048 | 54,860 | 55,672 | | | |
| 60,001-70,000 | 65, 805 | 66,766 | 67,727 | | | |
| 70,001-80,000 | 76,241 | 77,010 | 77,779 | | | |
| 80,001- 90,000 | 84, 115 | 85, 152 | 86, 189 | | | |
| 90, 001-100, 000 | 92, 942 | 94, 048 | 95, 154 | | | |
| 00, 001-120, 000 | 107,960 | 109, 462 | 110,964 | | | |
| 20, 001-140, 000 | 126, 144 | 128, 157 | 130, 170 | | | |
| 140, 001-160, 000 | 148,754 | 150, 926 | 153,098 | | | |
| 160, 001-200, 000 | 173, 110 | 178, 395 | 183,680 | | | |
| 200, 001-250, 000 | 215, 937 | 225, 667 | 235, 397 | | | |
| 50, 001-300, 000 | 260, 331 | 268, 949 | 277, 567 | | | |
| 100, 001-400, 000 | 331, 395 | 355, 375 | 379, 355 | | | |
| 100, 001-500, 000 | 402, 540 | 448, 863 | 495, 186 | | | |
| 00, 001 and over | 855, 882 | 955, 667 | 1,055,452 | | | |

Table 2.5

Projected Number $R_t^{(h)} \stackrel{1}{=} '$ of Individual Taxable Returns, for 1964 and 1965 and Estimates of the Mean Accumulated Taxes a_h

| Bracket | Per Cent | R ₄ ^(h) | R ₅ ^(h) | | a _h | |
|---------------|-----------------|-------------------------------|-------------------------------|--------------|----------------|-----------|
| ousand Pesos) | Distribution | 1963 | 1964 | Low | Middle | High |
| 0 - 2 | 63.58 | 96,042 | 107, 555 | 18.81 | 20.91 | 23.0 |
| 2 - 4 | 15.421 | 23, 293 | 26, 086 | 112.02 | 118.14 | 124.2 |
| 4 - 6 | 6.905 | 10,430 | 11,680 | 314.73 | 326.97 | 339.2 |
| 6 - 8 | 3.764 | 5,686 | 6,367 | 571.36 | 594.24 | 617.1 |
| 8 - 10 | 2.390 | 3,610 | 4,043 | 984.80 | 1,019.80 | 1, 054.8 |
| 10 - 20 | 4.345 | 6,563 | 7,350 | 1,815.12 | 2,039.76 | 2, 264.4 |
| 20 - 30 | 1.568 | 2,368 | 2,652 | 4,647.90 | 4,916.20 | 5, 182. 5 |
| 30 - 40 | .726 | 1,097 | 1, 228 | 7,797.60 | 8, 123.76 | 8, 449.9 |
| 40 - 50 | . 426 | 643 | 721 | 12, 421.60 | 12,755.60 | 13, 089.6 |
| 30 - SO | . 263 | 379 | 445 | 15, 780. 16 | 16, 121. 20 | 16, 462. |
| 60 - 70 | .140 | 211 | 237 | 20, 834. 20 | 21, 257. 04 | 21,678.8 |
| 70 - 80 | .091 | 137 | 154 | 25, 550.86 | 25,904.60 | 26, 259.3 |
| 80 - 90 | .063 | 95 | 107 | 29, 255.20 | 29, 752. 96 | 30, 250. |
| 90 -100 | .061 | 92 | 103 | 33,551.00 | 34, 104.00 | 34,657. |
| 100-120 | . 077 | 116 | 130 | 41, 219. 20 | 42,000.24 | 42,781. |
| 120-140 | . 049 | 74 | 83 | 50,736.32 | 51,803.21 | 52, 870. |
| 140-160 | . 026 | 39 | 44 | 62,807.16 | 63, 980. 04 | 65, 152. |
| 160-200 | . 036 | 54 | 61 | 76,090.50 | 78,997.25 | 81, 904. |
| 200-250 | . 017 | 26 | 29 | 99, 804.72 | 105, 253. 52 | 110, 702. |
| 250-300 | .012 | 18 | 20 | 124,768.67 | 129,680.93 | 134, 593. |
| 300-400 | .007 | 11 | 12 | 165, 589. 10 | 179, 497.50 | 193.405.9 |
| 400-500 | .006 | 9 | 10 | 206,878.60 | 234, 209.17 | 261.539. |
| Above 500 | .011 100.000 | 17 151, 028 | 19 169, 136 | 477, 909. 20 | 537, 780, 20 | 597,651. |

The projection of R_t is somewhat higher than the projection using simple regression without transforming R_t into its logarithm.

(C) Denote by \overline{x}_h the mean taxable assessment for the hth bracket. If r_h is the marginal rate for the hth bracket, an estimate of the nean accumulated tax a_h is

(5)
$$a_h = \sum_{j=1}^{h-1} r_j (b_{j2} - b_{j1}) + r_h (\overline{x}_h - b_{h-1})$$

where b_{j1} and b_{j2} are, respectively, the "lower" and "upper" boundaries of the jth bracket. Total tax assessment (middle) of the hth bracket may now be estimated by multiplying the projected number $R_t^{(h)}$ in year t of the hth bracket obtained from Table 2.5 with equation (5) above.

Since the variance of
$$a_h$$
 is
$$V(a_h) = r_h^2 V(\overline{x}_h),$$

the standard error of a_h may be computed from \underline{low} and \underline{high} estimates of \overline{x}_h given in Table 2.4 by simply multiplying 1/2 the difference of \underline{high} and \underline{low} estimates by the corresponding r_h . Using (5) and the standard error of a_h , \underline{low} and \underline{high} estimates of a_h can be computed from the 95% confidence limits for a_h . These are also given in Table 2.5. From (B) and (C) above, Table 2.5 may now be utilized to obtain the projected \underline{low} , \underline{middle} and \underline{high} individual tax assessment by multiplying $R_t^{(h)}$ by the corresponding \underline{low} , \underline{middle} and \underline{high} estimates for a_h . The results are tabulated (Table 2.6) below:

Table 2.6

PROJECTIONS OF TOTAL INDIVIDUAL TAX ASSESSMENTS BY BRACKETS: 1963-1964 (Pesos)

| | : | | 1963 | | : | | 1964 | |
|---|--------|------------------------|------------------------|-------------|---|------------|-------------|-------------|
| Brac | kets: | Low | Middle | High | : | Low | Middle | High |
| 0 - 2 | 2,000: | 1,806,550 | 2,008,238 | 2,209,926 | : | 2,023,110 | 2,248,975 | 2,474,84 |
| | 4,000: | 2,609,282 | 2,751,835 | 2,894,388 | : | 2,922,154 | 3,081,800 | 3,241, |
| 001 - 6 | | 3,282,634 | 3,410,297 | 3,537,960 | : | 3,676,046 | 3,819,010 | 3,961,973 |
| Control of | 8,000: | 3,248,753 | 3,378,849 | 3,508,944 | : | 3,637,849 | 3,783,526 | 3,929,203 |
| | 0,000: | 3,555,128 | 3,681,478 | 3,806,745 | : | 3,981,546 | 4,123,051 | 4, 264, 55 |
| | 0,000: | 11,912,633 | 13,386,945 | 14,861,257 | : | 13,341,132 | 14,992,236 | 16,643,34 |
| | :000,0 | 11,006,227 | 11,639,194 | 12,272,160 | : | 12,326,231 | 13,035,110 | 13,743,99 |
| | 0,000: | 8,553,967 | 8,911,765 | 9,269,562 | | 9,575,453 | 9,975,977 | 10,376,50 |
| | 0,000: | 7,987,089 | 8,201,851 | 8,416,613 | | 8,955,974 | 9,196,788 | 9,437,60 |
| | 0,000: | 6,264,724 | 6,400,116 | | : | 7,022,171 | 7,173,934 | 7,325,69 |
| | 0,000: | 4,396,016 | 4,485,235 | 4,574,244 | | 4,937,705 | 5,037,918 | 5, 137, 89 |
| | 0,0003 | 3,500,468 | 3,548,930 | 3,597,393 | | 3,934,832 | 3,989,308 | 4,043,78 |
| | 0,000: | 2,779,244 | 2,826,531 | 2,873,818 | | 3,130,306 | 3,183,567 | 3,236,82 |
| | 0,000: | 3,086,692 | 3,137,568 | 3,188,444 | | 3,455,753 | 3,512,712 | 3,569,67 |
| | 0,000: | 4,781,427 | 4,872,028 | 4,962,628 | | 5,358,496 | 5,460,031 | |
| | 0,000: | 3,754,488 | 3,833,438 | 3,912,387 | | | | 5,561,56 |
| | 0,000: | 2,449,479 | | | | 4,211,115 | 4,299,666 | 4,388,21 |
| BOOK STATE OF THE PARTY OF THE | 0,000: | 4, 108, 887 | 2,495,222 | 2,540,964 | | 2,763,515 | 2,815,122 | 2,866,72 |
| | 0,000: | 2,594,923 | 4,265,852 2,736,592 | 4,422,816 | | 4,641,521 | 4,818,832 | 4,996,14 |
| | 0,000: | | | 2,878,260 | | 2,894,337 | 3,052,352 | 3,210,36 |
| | 0,000: | 2,245,836 | 2,334,257 | 2,422,677 | : | 2,495,373 | 2,593,619 | 2,691,86 |
| | 0,000: | 1,821,480 | 1,974,473 | 2,127,465 | | 1,987,069 | 2,153,970 | 2,320,87 |
| | 0,000: | 1,861,907 8,124,456 | 2,107,883 | 2,353,858 | | 2,068,786 | 2,342,092 | 2,615,39 |
| (V) 50 | | | 9,142,263 | 10,160,070 | • | 9,080,275 | 10,217,824 | 11,355,37 |
| 1 | | 105,732,290 | 111,530,840 | 117,328,088 | .1 | 18,420,749 | 124,907,420 | 131,373,03. |

2.4 Mean Taxable Income and Percent Distribution of Smaller Brackets

The sub-sample from the Manila list may be utilized to obtain the breakdown into smaller brackets. The proportions of smaller brackets in the larger brackets have been derived from the sample together with the mean taxable income in each smaller bracket. If the marginal rates are known for each of these sub-brackets average cumulated tax assessments may be made for each of the sub-brackets just as in Table 2.5. Subdividing the projected number of returns of the larger brackets by the known proportion of the smaller brackets (Table 2.3A) and multiplying these numbers by the average cumulated tax assessments will give us again the required tax assessments by smaller brackets.

Table 2.7 below shows the mean taxable incomes in the sub-brackets obtained from the 1,073 sub-sample of the Manila list. It may be remarked that, although proportions of married (or single) taxable returns cannot be used by pooling sub-samples of single and married groups (first paragraph of section 2.2A), the mean taxable incomes of sub-brackets may be pooled from the means of the sub-bracket in the single and married groups. Denote by $\overline{\mathbf{x}}_1$ the sample mean of single taxable return in a sub-bracket and $\overline{\mathbf{x}}_2$ the sample mean of married taxable return in the same sub-bracket. The mean taxable return $\overline{\mathbf{x}}$ for the particular sub-bracket is

$$\bar{x} = \frac{n_1'\bar{x}_1 + n_2\bar{x}_2}{n_1' + n_2'},$$

Table 2.7 PROPORTION AND MEAN TAXABLE INCOMES BY SMALLER BRACKETS

| Bracket | Sub-bracket | Proportion | Mean Taxable Income |
|--|----------------------------------|------------|---------------------|
| (P1,000) | (P1,000) | | (P1,000) |
| 0 - 2 | | 1:00 | |
| 0 - 2 | 0 - 1 | .6537 | .4524 |
| | 0 - 1 1 - 2 | .3463 | 1.4944 |
| | 1 - 2 | •3403 | 1,4744 |
| | | | 1 |
| 2 - 4 | | 1.00 | |
| Constitutive Constitution of C | 2 - 3 | .6207 | 2.4444 |
| | 3 - 4 | .3793 | 3.4545 |
| | | | |
| | | | |
| 4 - 6 | | 1.00 | |
| | 4 - 5 | .5152 | 4.5882 |
| | 5 - 6 | . 4848 | 5.5000 |
| | | | |
| | | | |
| 6 - 8 | | 1.00 | |
| | 6 - 7 7 - 8 | .5636 | 6.5806 |
| | 7 - 8 | .4364 | 7.5000 |
| | | | |
| | | | |
| 8 - 10 | | 1.00 | |
| | 8 - 9 | .6667 | 8.5357 |
| | 9 - 10 | .3333 | 9,5714 |
| | | | |
| 10 - 20 | | 1.00 | |
| | 10 - 12 | .3846 | 10.8000 |
| | 12 - 14 | .2051 | 13.0000 |
| | 14 - 16 | .2051 | 15,0000 |
| | 16 - 18 | .0769 | 17.6667 |
| | 18 - 20 | .1282 | 19.0000 |
| | | | |
| | | | |
| 20 - 30 | | 1.00 | |
| | 20 22 | .2195 | 20.7778 |
| | 22 - 24 | .1951 | 23,1250 |
| | 24 - 26 | .2195 | 24,8889 |
| | 26 - 28 | .1951 | 26,7500 |
| | 28 - 30 | .1707 | 29.0000 |
| | | | |
| 20 40 | | .1 00 | |
| 30 - 40 | 30 - 32 | 1.00 | 31 3000 |
| | | . 2903 | 31,3000 |
| | 32 - 34 | . 16215 | 32.8333 |
| | 34 - 36 | .2703 | 34.8000 |
| | 36 - 38 38 - 40 | .1351 | 37,2000 |
| | 30 - 40 | .16215 | 38,8333 |

| | 200-0712 | TIENNIA SEC | A Ser Transfer Losso |
|------------------|----------------------|----------------|----------------------|
| Bracket (P1,000) | Sub-bracket (P1,000) | Proportion | Mean Taxable Income |
| (F1,000) | (F1,000) | | (P1,000) |
| 40 - 50 | | 1.00 | |
| | 40 - 42 | .2000 | 41.1111 |
| | 42 - 44 | .1333 | 43.1667 |
| | 44 - 46 | .2667 | 45.0000 |
| | 46 - 48 | .1556 | 46.7143 |
| | 48 - 50 | .2444 | 48.9091 |
| | | | |
| EO (O | | 1 00 | |
| 50 - 60 | 50 - 52 | 1.00 | 50.8125 |
| | 52 - 54 | . 2759 | 53.0833 |
| | 54 - 56 | .2069 .1552 | 54.7778 |
| | 56 - 58 | .1379 | 56.8750 |
| | 58 - 60 | .2241 | 59.0000 |
| | 36 - 00 | • 2241 | 39.0000 |
| | | | |
| 60 - 70 | | 1.00 | |
| | 60 - 62 | .2558 | 60.6360 |
| | 62 - 64 | .2093 | 63.0000 |
| | 64 - 66 | .1395 | 64.6670 |
| | 66 - 68 | .1860 | 67.1250 |
| | 68 - 70 | .2093 | 69.0000 |
| | | | |
| | | | |
| 70 - 80 | | 1.00 | |
| | 70 - 75 | .5079 | 72.4370 |
| | 75 - 80 | .4921 | 77.8710 |
| | | | |
| 80 - 90 | | 1.00 | |
| 00 = 50 | 80 - 85 | .4848 | 82,4370 |
| | 85 - 90 | .5152 | 87.9410 |
| | 03 - 70 | .5152 | 0 |
| | | | |
| 90 -100 | | 1.00 | |
| | 90 - 95 | .5556 | 92.1330 |
| | 95 -100 | . 4444 | 97.2500 |
| | | | |
| 100 100 | | 1.00 | |
| 100-120 | 100 105 | 1.00 | 102 2520 |
| | 100-105 | .3400 | 102,3530 |
| | 105-110 | .2600 | 107.7690 114.0500 |
| | 110-120 | •4000 | 114,0300 |
| 120-140 | | 1.00 | |
| 150-170 | 120-130 | .6333 | 125.9470 |
| | 130-140 | .3667 | 135.7270 |
| | | | |

| Bracket (P1,000) | Sub-bracket (P1,000) | Proportion | Mean Taxable Income (P1,000) |
|--|----------------------|------------|------------------------------|
| 140-160 | | 1.00 | |
| mediana di manana di | 140-145 | .5833 | 142.1430 |
| | 145-150 | .2500 | 147.3330 |
| | 150-155 | .1670 | 151.0000 |
| | 155-160 | . 0 | |
| | | | |
| 160-200 | | 1.00 | 145 0000 |
| | 160-170 | .3333 | 165.0000 |
| | 170-180 | .3333 | 175.1429 |
| | 180-190 | .1111 | 185.0000 |
| | 190-200 | .2223 | 196.4000 |
| 200-250 | | 1.00 | |
| The state of the s | 200-210 | .2500 | 204,6670 |
| | 210-220 | .0833 | 214.0000 |
| | 220-230 | .0833 | 226,0000 |
| | 230-240 | .2500 | 236.3333 |
| | 240-250 | .3300 | 245,0000 |
| | | | |
| 250-300 | | 1.00 | |
| | 250-260 | .3636 | 253.2500 |
| | 260-270 | .0909 | 265.0000 |
| | 270-280 | .2727 | 274.3333 |
| | 280-290 | .1818 | 281.0000 |
| | 290-300 | .0909 | 296,0000 |
| | | | |
| 300-400 | | 1.00 | |
| | 300-350 | .800 | 319.0000 |
| | 350-400 | .200 | 376.0000 |
| 400-500 | | 1.00 | |
| 400=300 | 400 450 | •6667 | 425,5000 |
| | 490-450 | | |
| | 450 –5 00 | .3333 | 496.0000 |
| Above 500 | | 1.00 | 907.143 |
| | | | |

where n_1' and n_2' are, respectively, the sample sizes of single and married groups in the same sub-bracket.

2.5 Projections Under Proposed Brackets (PIA).

Following the brackets proposed by PIA, individual tax assessments for 1963 and 1964 have been worked out (Table 2.8). The number of returns in the proposed brackets were each derived from the sample proportions shown in Table 2.7 with some adjustments if the proposed ones have boundaries other than those obtainable in that table. The mean cumulated tax assessments were similarly derived as in section 2.3 (equation 5) under the proposed marginal rates. The mean taxable incomes were obtained from the sample data. Compared with Table 2.6 the middle estimates for the corresponding years and the projections total under the PIA proposal differ by about \$11\$ million.

2.6 The Projection of Total Individual Income Tax

The projections in section 2.3 of the individual income tax assessments by brackets for the years 1963 and 1964 are cross-sectional in nature. The total of middle estimates are values resulting from an estimator of the total individual tax assessments. For some fixed number of individual taxable returns $R_t^{(h)}$ in year t of the $h^{\underline{th}}$ bracket the tax assessment is $R_t^{(h)}$ x a_h , where a_h is now the average accumulated tax assessment in the $h^{\underline{th}}$ bracket. Since a_h (from (5) of that section) is a linear function of \overline{x}_h (mean taxable income in the $h^{\underline{th}}$ bracket) and \overline{x}_h in large